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Genomewide Association, Parkinson Disease, and *PARK10*

To the Editor:

Genomewide linkage analysis of rare familial forms of parkinsonism has identified mutations in seven genes, revealing a clinicopathologically and genetically heterogeneous syndrome.¹ Less progress has been made in the more typical late-onset form of Parkinson disease (PD [MIM 168600]), although the recently identified LRRK2 (MIM 609007) G2019S substitution is estimated to account for ~1% of sporadic PD cases.² Common polymorphisms of familial genes may also influence susceptibility to idiopathic PD.^{3,4} Of the 198,345 SNPs successfully genotyped in the recent genomewide association (GWA) study, 26 had notably different allele frequencies between patients and controls in both tiers (P < .01).⁵ Fifteen of these SNPs had opposite directions of effect (disease risk or protection) in tiers 1 and 2. The remaining 11 SNPs were proposed as markers for new genes/chromosomal loci that influence susceptibility to PD. In addition, two SNPs in tier 2 (rs682705 and rs7520966) were highlighted in the PARK10 locus (MIM 606852), which nominated the gene LOC200008 in disease susceptibility.

The PARK10 locus on chromosome 1p32 was originally identified in a genomewide linkage analysis of 117 patients from 51 Icelandic families (maximum $Z_{\rm lr}$ = 4.8 at D1S231, with a LOD-1, 7.6-cM support interval from D1S2874 to D1S475).6 Iceland has a well-characterized genealogy that is powerful for family-based linkage studies. The ancestral founders of Iceland have Scandinavian patrilineal inheritance with a minor Celtic matrilineal component.⁷ Assuming that the PARK10 mutation predates the Icelandic settlement, we reasoned that the 1p32 susceptibility gene might be more readily found in patients with PD originating from Scandinavian or Celtic populations. In parallel to the study of Maraganore et al.,5 we have been mapping the PARK10 locus. Genotypes from 28 SNPs (including rs682705 and rs7520966) within a 132-kb region of chromosome 1p32 located around the LOC200008 gene have been analyzed in two well-characterized case-control series from Norway and Ireland. In addition, we attempted to replicate findings for the two *PARK10* SNPs in a U.S. series collected at the Mayo Clinic in Jacksonville, FL. We then employed all three case-control series to investigate the genotype/allele frequencies of the main 11 SNPs nominated to influence PD susceptibility.⁵ Power was comparable to the original study (>80% at α = 0.05 for odds ratios [ORs] >2.0 and for disease-allele frequencies >0.035), and genotyping call rates were >95% for all markers (table 1).

In total, Norwegian samples included 676 subjects (cases and controls) with a mean age (\pm SD) of 70 \pm 11 years, Irish samples included 372 subjects with a mean age (\pm SD) of 61 \pm 13 years, and the U.S. samples included 522 subjects with a mean age (\pm SD) of 71 \pm 10 years. All patients were examined and were observed longitudinally by a movement-disorders neurologist (J.O.A., J.M.G., D.G., T.L., Z.K.W., and R.J.U.), and they were given a diagnosis of PD in accordance with published criteria.⁸ Each patient was individually matched, on the basis of age (\pm 4 years) and ethnicity, to an unrelated control without evidence of neurological disease. The ethical review boards at each institution involved approved the study, and all participants provided informed consent.

SNP genotyping was performed using TaqMan chemistry on an ABI7900 genetic analyzer; in cases where genotype data was available for only one subject of a matched pair, the other subject was retained in the analysis. For the controls in each population, χ^2 tests of Hardy-Weinberg equilibrium (HWE) were implemented using Haploview.⁹ Optimal SNP coverage for association analysis of the *LOC200008* gene was determined empirically by the construction of linkage-disequilibrium (LD) maps in Norwegian and Irish samples, onto which haplotype blocks were assigned (fig. 1).^{10,11} ORs for disease association, with corresponding 95% CIs, were subsequently calculated using logistic-regression models adjusted for age and sex. Overall ORs combining data from

Table 1

Genotype/Allele Frequencies of the Main 11 SNPs Nominated to Influence PD Susceptibility

The table is available in its entirety in the online edition of *The American Journal of Human Genetics*.



Figure 1 Metric LD map and haplotype block structure of the investigated region. *A*, LD map providing information about LD patterns in the investigated candidate region, through locations expressed in LDUs. LDUs have an inverse relationship with LD, with regions of extensive recombination having many LDUs. The physical position of the gene in the region LOC200008 is marked with an arrow. All 28 SNPs genotyped are reported, although the symbols (\blacktriangle) may be obscured for SNPs that lie in close physical proximity and high LD. SNPs *rs682705* and *rs7520966* are denoted by an asterisk (*). *B*, LD structure of the candidate region. *Black* and *dark gray cells*, strong LD; *gray cells*, intermediate; and *light gray* and *white cells*, evidence for historical recombination. The haplotype block structure of the region is defined according to Gabriel et al.¹⁰ An asterisk denotes SNPs *rs682705* and *rs7520966*. The LD map and haplotype structure were constructed using genotypes from the Norwegian sample. Similar results were obtained for the Irish population.

all three sites were additionally adjusted for site. Previous studies have nominated the *PARK10* locus as an age-at-onset modifier in PD¹²; thus, we also assessed the influence of 1p32 SNPs variability on this disease trait, using linear-regression models adjusted for sex.

There was no evidence of association with PD for any of the 28 genotyped 1p32 SNPs in our study (all SNP P > .05 after applying Bonferroni correction in both population samples). Haplotype frequencies between patients and controls were not significantly different for the haplotype blocks identified; nor was the age at onset in patients associated with any single marker or haplotype (all corrected P > .05). Of note, the ancestral recombination and haplotype blocks apparent within Norwegian and Irish samples were comparable for this interval at this marker resolution. The average number of SNPs per LD unit (LDU) was 6.8 (mean LDU between markers 0.15, range 0–0.63), indicating that the number of SNPs genotyped within and flanking *LOC200008* should be sufficient for examination of the region.¹¹ In addition, the two *PARK10* SNPs showed no significant association within the U.S. series (P > .05). None of the other 11 SNPs nominated by the GWA study had different allele frequencies or genotype distributions between affected subjects and matched controls (all SNP P > .05 in all populations independently or as a combined sample set) (table 2). There was no evidence of departure from HWE in controls (P > .01 in all population controls).

Our study indicates that genetic variability within the LOC200008 gene is unlikely to explain the PARK10 susceptibility locus for PD. Sadly, the lack of disease association and replication in an independent U.S. series of comparable power suggests that the original findings may be spurious. Failure to nominate LOC200008 as the PARK10 gene in our population samples provides empirical support for statistical caveats concerning GWA studies. Implicit in multiple testing is false discovery, even in well-designed studies, and there are several potential sources of bias.¹³ Of note, neither PARK10 SNP rs682705 nor rs7520966 fulfilled the main criterion for being genotyped in tier 2 (P < .01 in tier 1 overall analvsis), but each was included with a less stringent association criterion (P < .05 in tier 1 overall analysis) because of its physical position within a PARK locus. Interestingly, the combined P value for rs682705 ($P = 9.07 \times 10^{-6}$) is the second-lowest P value of the overall study, even though it did not fulfill the inclusion criteria. Individuallevel data from the GWA study is not yet available, but, in our study, these two SNPs also appear to be in LD (pairwise $r^2 > 0.9$), as suggested by Maraganore et al.⁵; in addition, the minor-allele frequencies (MAFs) of the two SNPs are comparable across studies and populations. The former suggests less-than-optimal haplotype tagging in the initial study, whereas the latter argues against technical errors in genotyping, but neither provides sufficient explanation for the positive findings observed elsewhere.5

We found no evidence of direct association between the 11 SNPs nominated in the GWA study and disease in the three independent populations or in a combined sample group (n = 1,570) (table 2). However, for these loci, we did not employ a gene-based approach (nor did we fine-map each region as with PARK10), as advocated elsewhere¹⁴; we await the results of further replication studies. Of note, in the study by Maraganore et al.,⁵ the rs7702187 SNP within SEMA5A (MIM 609297) had the lowest combined P value ($P = 7.62 \times 10^{-6}$); however, a total of 53 SNPs were examined in this gene in tier 1. Only rs7702187 was significant before correction (P = .001), which supports the possibly spurious nature of this and the other associations. The MAFs observed in our three populations and in that of the GWA study are comparable, which argues against population bias/ heterogeneity (table 2).

The number of SNPs highlighted in each tier of the original study is consistent with chance—that is, 1% of SNPs use a significance level of P < .01. None of the P values obtained by Maraganore et al.⁵ meets a Bonferroni correction for multiple testing, although this standard may be too conservative in GWA, since it fails to account for LD and incorrectly assumes that chromosomal markers are independent. A consensus on the most appropriate correction for multiple testing has yet to be reached. Now that genomewide data sets have been generated, there exists the possibility to use these to develop appropriate statistical methods to identify true positive results.¹⁵

In the interim, we recommend that enthusiasm for positive findings should be tempered by the strength of the evidence, the population-attributable risk, and the differences in SNP allele/genotype frequencies between cases and controls. If allele frequencies are significantly different, genomic controls might be used to assess population substructure. It is important that future studies

Table 2

11 SNPs Nominated in GWA Study	as Genetic Susceptibility Loci for PD
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	Chromosome	Position (bp)	This Study				MARAGANORE ET AL. ⁵						
dbSNP Accession			Control MAF		Estimated OR	Combined P	Control MAF		Р		Estimated OR	Combined P	
NUMBER			Ireland	Norway	United States	(95% CI) ^a	(n = 1,570)	Tier 1	Tier 2	Tier 1	Tier 2	(95% CI) ^b	(n = 1,550)
rs7702187	5p15.2	9385281	.17	.18	.18	.88 (.74-1.06)	.18	.18	.20	.001	.002	1.74 (1.36-2.24)	7.62×10^{-6}
rs10200894	2q36	228642637	.13	.11	.09	.96 (.77-1.21)	.74	.12	.13	.009	.001	1.84 (1.38-2.45)	1.70×10^{-5}
rs2313982	4q31.1	139145665	.05	.11	.09	.93 (.73-1.18)	.54	.07	.06	.006	.002	2.01 (1.44-2.79)	1.79×10^{-5}
rs17329669	7p14	36625169	.13	.12	.14	1.01 (.82-1.24)	.92	.13	.11	.008	.001	1.71 (1.33-2.21)	2.30×10^{-5}
rs7723605	5p15.3	5407615	.13	.14	.13	.91 (.75-1.12)	.38	.12	.09	.010	.002	1.78 (1.35-2.35)	3.30×10^{-5}
ss46548856	10q21	58986929	.09	.08	.11	.93 (.73-1.19)	.58	.09	.11	.003	.002	1.88 (1.38-2.57)	3.65×10^{-5}
rs16851009	2q24	166456214	.11	.11	.12	.95 (.76-1.18)	.63	.09	.08	.002	.009	1.84 (1.36-2.49)	4.17×10^{-5}
rs2245218	1p36.2	13885132	.19	.17	.15	.95 (.79-1.14)	.57	.11	.13	.002	.002	1.67 (1.29-2.14)	4.61×10^{-5}
rs7878232	Xq28	150516943	.25	.23	.25	1.10 (.97-1.25)	.15	.29	.26	.003	.010	1.38 (1.17-1.62)	6.87×10^{-5}
rs1509269	4q31.1	139111329	.08	.13	.13	.94 (.76-1.17)	.58	.10	.09	.005	.008	1.71 (1.30-2.26)	9.21×10^{-5}
rs11737074	4q27	125438978	.21	.20	.21	1.05 (.89-1.25)	.55	.19	.19	.007	.005	1.50 (1.21-1.86)	1.55×10^{-4}

NOTE.—In this study, MAFs are not significantly different between the populations. No P values are corrected for multiple testing. SNPs are ordered by combined P value, per Maraganore et al.⁵

^a The direction of effect of the estimated OR observed in this study for each SNP is shown (i.e., >1 risk and <1 protective).

^b Estimated ORs in the study by Maraganore et al.⁵ do not indicate the direction of effect relative to the MAF.

employ multiple independent sample series, each with sufficient power to verify significant genetic associations, before publication.¹⁶ However, lack of evidence for an association is not the same as evidence against one; thus, lack of replication should also be interpreted with caution.

Over the few next years, the number of GWA studies will increase, and it is important to learn from the experiences gained by the few studies performed to date. Although our negative findings suggest that the conclusions drawn from the study by Maraganore et al.⁵ might be based on spurious associations, further analysis of individual-level raw data is now necessary. The recent identification of a complement factor H polymorphism in age-related macular degeneration in a GWA study and the identical findings by two other groups using other study designs demonstrates that this approach can be used successfully.¹⁷⁻¹⁹ It may be that, because of the heterogeneous nature of PD, associations with a gestalt phenotype are masked by background variation in SNP informativeness, population strata, and insufficient power. It is, therefore, crucial that future associations are validated and that analysis is performed to resolve the underlying cause of association in the sample population. GWA studies may still provide direction for the genetic analysis of heterogenous complex traits, but, in the short term, they may exacerbate the problem of replication failure in association studies.

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Web Resources

The URLs for data presented herein are as follows:

- dbSNP, http://www.ncbi.nlm.nih.gov/entrez/query.fcgi?CMD = search &DB = snp
- Online Mendelian Inheritance in Man (OMIM), http://www.ncbi.nlm .nih.gov/Omim/ (for PD, *LRRK2*, *PARK10*, and *SEMA5A*)

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